

Bayesian filters for continuous-discrete dynamic models

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STOCHASTICA Directions Workshop
Cork, 29 April 2026.

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Partially supported by grant PID2024-158181NB-I00 NISA, funded by MCIN/AEI/10.13039/501100011033 and ERDF, and grant IDEA-CM TEC-2024/COM-89, funded by Community of Madrid. Part of this research was performed at the mathematical research institute MATRIX.

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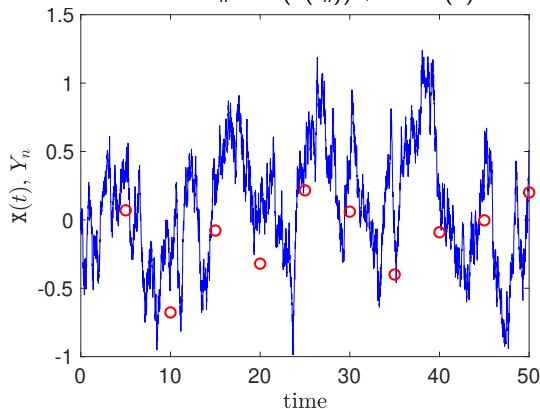
Barrier functions

Discussion

References

Continuous-discrete state space model

- ▶ Continuous-time stochastic state $X(t)$ (—)
- ▶ Discrete-time observations $Y_n = m(X(t_n)) + \text{noise}$ (○)



- ▶ The aim is to compute $\mathbb{P}(X(t) \in dx \mid \{Y_n, n \in \mathbb{N} : t_n \leq t\})$.

Continuous-discrete state space models

- Assume the signal of interest is an Itô process $X(t)$ in a probability space $(\Omega, \mathcal{F}, \mathbb{P})$. The dynamics are described by the SDE

$$dX(t) = \underbrace{a(X(t), t)}_{\text{drift}} dt + \underbrace{\sigma(X(t), t)}_{\text{diffusion coeff.}} \underbrace{dW(t)}_{\text{Wiener proc.}}, \quad t \in [0, T].$$

- For a given time grid $0 = t_0 < t_1 < \dots < t_n < \dots < t_N = T$:
 - Let $X_n = X(t_n)$, the observation at time t_n is

$$Y_n = m(X_n) + U_n, \quad n = 1, 2, \dots$$

where U_n is iid noise and $m : \mathcal{X} \mapsto \mathcal{Y}$ is an observation function.

- The sequence $X_n := X(t_n)$, $n \geq 0$, is described by the Markov kernels

$$K_n(x', dx) := \mathbb{P}(X_n \in dx | X_{n-1} = x'), \quad \text{for } n \geq 1.$$

- The likelihood at time t_n , denoted $g_n(x) \propto \bar{g}(y_n - m(x))$, depends on the law $\bar{g}(u)$ of the noise variables U_n and the observed value $Y_n = y_n$.

Example: if $U_n \sim \mathcal{N}(0, \Sigma_u) \implies g_n(x) \propto \exp \left\{ -\frac{1}{2} \|y_n - m(x)\|_{\Sigma_u^{-1}}^2 \right\}$

Bayesian filtering

- ▶ Given the kernels K_n and likelihoods g_n , we are interested in the marginal probability laws

$$\underbrace{\xi_n(dx) := \mathbb{P}(X_n \in dx | Y_{1:n-1} = y_{1:n-1})}_{\text{one-step-ahead prediction}}, \quad \underbrace{\pi_n(dx) := \mathbb{P}(X_n \in dx | Y_{1:n} = y_{1:n})}_{\text{Bayesian filter}}$$

- ▶ Recursive construction

$$\text{(predict)} \quad \xi_n(dx) = K_n \pi_{n-1}(dx) := \int_{\mathcal{X}} K_n(x', dx) \pi_{n-1}(dx'),$$

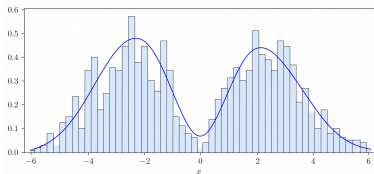
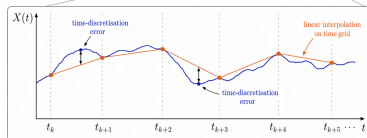
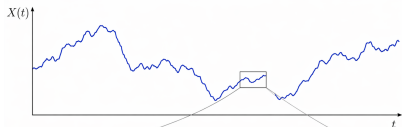
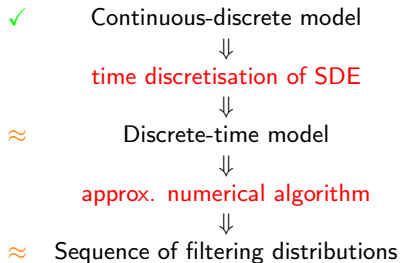
$$\text{(update)} \quad \pi_n(dx) = \frac{g_n(x) \xi_n(dx)}{\xi_n(g_n)}, \quad \text{where } \xi_n(g_n) = \int_{\mathcal{X}} g_n(x) \xi_n(dx)$$

- ▶ Issues: in general...
 - we can only sample K_n (approximately),
 - both integrals are intractable.

Need for efficient numerical methods!

Approximation errors

Approximation pipeline



Time-discretisation errors

- ▶ The kernel K_n is intractable. We can (approximately) sample, e.g.,

$$\bar{X}_{n,0}^h = x', \quad \text{and}$$

$$\bar{X}_{n,j}^h = \bar{X}_{n,j-1}^h + ha(\bar{X}_{n,j-1}^h, t'_{n,j-1}) + \sqrt{h}\sigma(\bar{X}_{n,j-1}^h, t'_{n,j-1})Z_{n,j},$$

for $j = 1, \dots, K$ and $h = t'_{n,j} - t'_{n,j-1}$. Then $X_n^h := \bar{X}_{n,J}^h \approx X_n = \mathbf{X}(t_n)$.

- ▶ True vs. approximate model: given observations $Y_n = y_n$,

continuous time			discrete time			
X_0	\sim	$\pi_0(dx)$		X_0^h	\sim	$\pi_0(dx)$
X_n	\sim	$K_n(x_{n-1}, dx)$,	X_n^h	\sim	$K_n^h(x_{n-1}, dx)$
$g_n(X_n)$	\propto	$\bar{g}(y_n - m(X_n))$		$g_n(X_n^h)$	\propto	$\bar{g}(y_n - m(X_n))$

A choice of numerical scheme yields an approximate kernel K^h . For a numerical scheme of weak order p

$$\sup_{0 \leq n \leq N} \left| \mathbb{E}[\phi(X_n)] - \mathbb{E}[\phi(X_n^h)] \right| < C_\phi h^p \quad \text{for a test function } \phi \in C_B^4(\mathbb{R}^{d_x}).$$

Time-discretisation errors

- ▶ Time-discretisation affects the posterior marginals

$$\underbrace{\begin{aligned} \xi_n(\mathrm{d}x) &= (K_n \pi_{n-1})(\mathrm{d}x) \\ \pi_n(\mathrm{d}x) &= \frac{g_n(x) \xi_n(\mathrm{d}x)}{\xi_n(g_n)} \end{aligned}}_{\text{continuous time}}, \quad \underbrace{\begin{aligned} \xi_n^h(\mathrm{d}x) &= (K_n^h \pi_{n-1}^h)(\mathrm{d}x) \\ \pi_n^h(\mathrm{d}x) &= \frac{g_n(x) \xi_n^h(\mathrm{d}x)}{\xi_n^h(g_n)} \end{aligned}}_{\text{discrete time}}$$

Theorem (Time-discretisation, Akyildiz *et al.* (2024); Erdogan *et al.* (2026))

Let $Y_{1:N} = y_{1:N}$ be arbitrary but fixed, let K_n^h be a weak-order 1 scheme and choose $\phi \in C_B^4(\mathbb{R}^{d_x})$ with $\|\phi\|_\infty \leq 1$. If

- ▶ $g_n > 0$, $\|g_n\|_\infty \leq 1$ and $g_n \in C_B^4(\mathbb{R}^{d_x})$, and
- ▶ $\bar{\phi}_n \in C_B^4(\mathbb{R}^{d_x})$, where $\bar{\phi}_n(x') = \int \phi(x) K_n(x', \mathrm{d}x)$,

then there are finite constants $\{\tilde{C}_n, C_n\}_{n=1}^N$, independent of h , such that

$$\left| \xi_n(\phi) - \xi_n^h(\phi) \right| \leq \tilde{C}_n h \quad \text{and} \quad \left| \pi_n(\phi) - \pi_n^h(\phi) \right| \leq C_n h \quad \text{for } n = 1, \dots, N.$$

Monte Carlo error example: the particle filter

- ▶ Numerical filtering methods add approximation error on top of the discretisation error.

Example: standard particle filter

- ▶ Algorithm

1. Draw $x_0^{h,i} \sim \pi_0$, $i = 1, \dots, M$.
2. Recursive step: for $i = 1, \dots, M$
 - i. draw $\bar{x}_n^{h,i} \sim K_n^h(x_{n-1}^{h,i}, dx)$,
 - ii. weight $w_n^i \propto g_n(\bar{x}_n^{h,i})$,
 - iii. resample $\{\bar{x}_n^{h,i}, w_n^i\}_{i=1}^M \rightarrow \{x_n^{h,i}\}_{i=1}^M$
3. Approximation $\pi_n^{h,M}(dx) = \frac{1}{M} \sum_{i=1}^M \delta_{x_n^{h,i}}(dx)$.

- ▶ Approximation error (Erdogan *et al.*, 2026): under smoothness conditions, there are constants $C_n, B_n < \infty$ such that

$$\mathbb{E} \left[\left| \pi_n(\phi) - \pi_n^{h,M}(\phi) \right|^p \right]^{\frac{1}{p}} < \underbrace{C_n h}_{\text{discretisation}} + \underbrace{\frac{B_n}{\sqrt{M}}}_{\text{Monte Carlo}}, \quad p \geq 1.$$

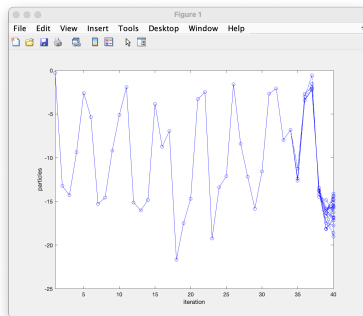
Modified SDEs & Monte Carlo filters

Particle filters

Generic particle filter with sampling kernel Q_n & weight degeneracy

1. Draw $x_0^i \sim \pi_0$, $i = 1, \dots, M$.
2. Recursive step: for $i = 1, \dots, M$
 - i. draw $\bar{x}_n^i \sim Q_n(x_{n-1}^i, dx)$,
 - ii. weight

$$w_n^i \propto g_n(\bar{x}_n^i) \frac{dK_n}{dQ_n}(x_{n-1}^i, \bar{x}_n^i),$$
 - iii. resample $\{\bar{x}_n^i, w_n^i\}_{i=1}^M \rightarrow \{x_n^i\}_{i=1}^M$



The “optimal” kernel is $Q_n(x_{n-1}^i, dx) = \mathbb{P}(X_n \in dx \mid X_{n-1} = x_{n-1}^i, Y_n = y_n)$.

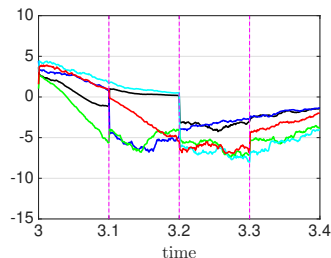
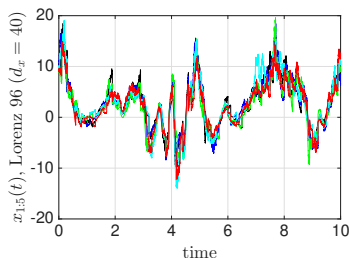
Weight degeneracy is an acute problem in high dimensional models!

- └ Modified SDEs for efficient Monte Carlo filtering

- └ Feedback particle filter

Feedback particle filter

- ▶ If weight degeneracy is a problem... remove the weights.
- ▶ Prediction using the SDE as in a standard PF.
Update maps predictive-law particles into filtering-law particles



- ▶ Conceptually appealing, but
 - “jumps” in sampled paths,
 - update only approximate (no exact sampling!)

Feedback particle filter (Yang *et al.*, 2014)

- ▶ Assume pdfs $\xi_n(x)$ (prediction) and $\pi_n(x)$ (filter).
How do we convert particle $\tilde{x}_n^i \sim \xi_n$ into particle $x_n^i \sim \pi_n$?
- ▶ Construct the homotopy of densities

$$\rho_n(x, \tau) \propto \xi_n(x) e^{\tau \Phi(x)}, \quad \text{where } \Phi(x) = m(x)^\top y_n - \frac{1}{2\sigma^2} \|m(x)\|^2$$

This yields $\rho_n(x, 0) = \xi_n(x)$ and $\rho_n(x, 1) \propto \pi_n(x)$, and

$$\partial_\tau \rho = \rho \left[(m(x) - (m, \rho))^\top y_n - \frac{1}{2\sigma^2} \|m(x)\|^2 + \frac{1}{2\sigma^2} (m^2, \rho) \right]$$

- ▶ Liouville (continuity) equation:

$$\text{if } \rho \left[(m(x) - (m, \rho))^\top y_n - \frac{1}{2\sigma^2} \|m(x)\|^2 + \frac{1}{2\sigma^2} (m^2, \rho) \right] = -\nabla \cdot (\rho U)$$

$$\text{then } \frac{d\tilde{x}}{d\tau} = U(\tilde{x}, \tau), \quad \text{with } \tilde{x}(0) = \tilde{x}_n^i, \quad \text{yields } \tilde{x}(1) = x_n^i \sim \pi_n$$

Finding $U(\tilde{x}, \tau)$ requires the numerical solution of a high dimensional Poisson equation \implies not easy to implement + approximation errors.

Feedback particle filter (Yang *et al.*, 2014)

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then $\frac{d\tilde{x}}{d\tau} = U(\tilde{x}, \tau)$, with $\tilde{x}(0) = \tilde{x}_n^i$, yields $\tilde{x}(1) = x_n^i \sim \pi_n$

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- └ Modified SDEs for efficient Monte Carlo filtering

- └ Feedback particle filter

Feedback particle filter (Yang *et al.*, 2014)

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Finding $U(\tilde{x}, \tau)$ requires the numerical solution of a high dimensional Poisson equation \implies not easy to implement + approximation errors.

- └ Modified SDEs for efficient Monte Carlo filtering

- └ Doob h -transform

The Doob h -transform (Rogers & Williams, 2000)

Recall the particle filter

1. Draw $x_0^i \sim \pi_0$, $i = 1, \dots, M$.
2. Recursive step: for $i = 1, \dots, M$
 - i. draw $\bar{x}_n^i \sim Q_n(x_{n-1}^i, dx)$,
 - ii. weight $w_n^i \propto g_n(\bar{x}_n^i) \frac{dK_n}{dQ_n}(x_{n-1}^i, \bar{x}_n^i)$,
 - iii. resample $\{\bar{x}_n^i, w_n^i\}_{i=1}^M \rightarrow \{x_n^i\}_{i=1}^M$

“Optimal” $Q_n(x_{n-1}^i, dx_n) := \mathbb{P}(X_n \in dx \mid X_{n-1} = x_{n-1}^i, Y_n = y_n)$.

- How to implement Q_n ? Modify the original SDE

$$d\mathbf{X}(t) = a(\mathbf{X}(t), t)dt + \sigma(\mathbf{X}(t), t)d\mathbf{W}(t)$$

$$\Downarrow$$

$$d\tilde{\mathbf{X}}(t) = [a(\tilde{\mathbf{X}}(t), t) + \sigma^2(\tilde{\mathbf{X}}(t), t)\nabla \log h(\tilde{\mathbf{X}}(t), t)] dt + \sigma(\tilde{\mathbf{X}}(t), t)d\mathbf{W}(t)$$

for $t \in (t_{n-1}, t_n]$, where $\sigma^2(\cdot, \cdot) = \sigma(\cdot, \cdot)\sigma(\cdot, \cdot)^\top$ and

$$h(x, t) := \mathbb{E}[g_n(\tilde{\mathbf{X}}(t_n)) \mid \mathbf{X}(t) = x]$$

is a Doob h -transform.

Particle filter via Doob h -transform

SDE for $t \in (t_{n-1}, t_n]$, with initial condition $X^i(t_{n-1}) = \tilde{X}^i(t_{n-1}) = x_{n-1}^i$

$$dX^i(t) = a(X^i(t), t)dt + \sigma(X^i(t), t)dW^i(t)$$

$$\Downarrow$$

$$d\tilde{X}^i(t) = [a(\tilde{X}^i(t), t) + \sigma^2(\tilde{X}^i(t), t)\nabla \log h(\tilde{X}^i(t), t)] dt + \sigma(\tilde{X}^i(t), t)dW^i(t)$$

- ▶ Let \mathbb{P}_t^i and $\tilde{\mathbb{P}}_t^i$ denote the laws of $X^i(t)$ and $\tilde{X}^i(t)$. Standard theory shows that, for $t_{n-1} < t \leq t_n$,

$$\frac{d\mathbb{P}_t^i}{d\tilde{\mathbb{P}}_t^i}(x) = \frac{h(X^i(t_{n-1}), t_{n-1})}{h(x, t)}, \quad \text{where } h(x, t) = \mathbb{E} [g_n(X^i(t_n)) \mid X(t) = x]$$

- ▶ Hence, at time $t = t_n$, we have

$$\frac{d\mathbb{P}_{t_n}^i}{d\tilde{\mathbb{P}}_{t_n}^i}(x) = \frac{\mathbb{E} [g_n(X^i(t_n)) \mid X^i(t_{n-1}) = x_{n-1}^i]}{g_n(x)}$$

and the factor $g_n(x)$ cancels in the weight eq. $w_n(x) \propto g_n(x) \frac{d\mathbb{P}_t^i}{d\tilde{\mathbb{P}}_t^i}(x)$,

$$w_n^i \propto \mathbb{E} [g_n(X^i(t_n)) \mid X^i(t_{n-1}) = x_{n-1}^i] \quad \text{independently of } \bar{x}_n^i := \tilde{X}^i(t_n)$$

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$$\Downarrow$$

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$$dX^i(t) = a(X^i(t), t)dt + \sigma(X^i(t), t)dW^i(t)$$



$$d\tilde{X}^i(t) = [a(\tilde{X}^i(t), t) + \sigma^2(\tilde{X}^i(t), t)\nabla \log h(\tilde{X}^i(t), t)] dt + \sigma(\tilde{X}^i(t), t)dW^i(t)$$

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- └ Modified SDEs for efficient Monte Carlo filtering

- └ Doob h -transform

Particle filter via Doob h -transform: summary

SDE for $t \in (t_{n-1}, t_n]$, with initial condition $X^i(t_{n-1}) = \tilde{X}^i(t_{n-1}) = x_{n-1}^i$

$$dX^i(t) = a(X^i(t), t)dt + \sigma(X^i(t), t)dW^i(t)$$

↓

$$d\tilde{X}^i(t) = [a(\tilde{X}^i(t), t) + \sigma^2(\tilde{X}^i(t), t)\nabla \log h(\tilde{X}^i(t), t)] dt + \sigma(\tilde{X}^i(t), t)dW^i(t)$$

Sampling, optimal kernel Q_n

- ▶ Sample $\bar{x}_n^i := \tilde{X}_n^i$ with initial condition $\tilde{X}^i(t_{n-1}) = x_{n-1}^i$
- ▶ Then $\bar{x}_n^i \sim \mathbb{P}(dX_n \mid X_{n-1} = x_{n-1}^i, Y_n = y_n)$, the “optimal” importance function
- ▶ The weight $w_n^i \propto \mathbb{E}[g_n(X^i(t_n)) \mid X^i(t_{n-1}) = x_{n-1}^i]$ is independent of \bar{x}_n^i .

Difficulties

Need to compute $\nabla \log h(\tilde{X}^i(t), t)$ and $\mathbb{E}[g_n(X^i(t_n)) \mid X^i(t_{n-1}) = x_{n-1}^i]$.
 See Chopin *et al.* (2023) for a neural network based implementation and Pieper-Sethmacher *et al.* (2025b,a) for linear approximations.

Constrained filters

- ▶ We have assumed so far that $\mathbf{X}(t) \in \mathcal{X} \subseteq \mathbb{R}^{d_x}$. The *truncation* or *constraint* of the state space \mathcal{X} by a sequence of compact subsets $\mathcal{C}_n \subset \mathcal{X}$ at times t_n has proved useful
 - algorithmically, to improve robustness (Garcia-Fernandez *et al.*, 2012; Wang *et al.*, 2020) and
 - theoretically, to improve stability (Crisan *et al.*, 2020; Erdogan *et al.*, 2025).
- ▶ A **constraint** is expressed by a sequence of compact subsets $\mathcal{C} = \{\mathcal{C}_n\}_{n \geq 0}$. The constraint modifies the prior law of X_0 and the Markov kernels

$$\hat{\pi}_0(d\mathbf{x}) = \frac{\mathbb{1}_{\mathcal{C}_0}(\mathbf{x})\pi_0(d\mathbf{x})}{\pi_0(\mathcal{C}_0)}, \quad \hat{K}_n(\mathbf{x}', d\mathbf{x}) = \frac{\mathbb{1}_{\mathcal{C}_n}(\mathbf{x})K_n(\mathbf{x}', d\mathbf{x})}{\int_{\mathcal{X}} \mathbb{1}_{\mathcal{C}_n}(\tilde{\mathbf{x}})K_n(\mathbf{x}', d\tilde{\mathbf{x}})}, \quad n \geq 1.$$

As a consequence, we obtain constrained predictions & filters

$$\hat{\xi}_n(d\mathbf{x}) = \hat{K}_n \hat{\pi}_{n-1}(d\mathbf{x}), \quad \hat{\pi}_n(d\mathbf{x}) = g_n(\mathbf{x}) \hat{\xi}_n(d\mathbf{x}) / \hat{\xi}_n(g_n), \quad n \geq 1.$$

How do we simulate $\hat{K}_n(\mathbf{x}', d\mathbf{x})$, a constrained SDE?

- Rejection sampling?
- Doob h -transform $h_n(\mathbf{x}, t) = \mathbb{P}(\mathbf{X}(t_n) \in \mathcal{C}_n \mid \mathbf{X}(t_{n-1}) = \mathbf{x}_{n-1}^i)$?

High-probability tubes

- ▶ Let us use Doob transforms to impose the constraints: for $t \in (t_{n-1}, t_n]$

$$dX^n(t) = [a(X^n(t), t) + \sigma^2(X^n(t), t) \nabla \log h_n(X^n(t), t)] dt + \sigma(X^n(t), t) dW(t),$$

where $h_n(x, t) := \mathbb{P}(X(t_n) \in \mathcal{C}_n \mid X(t) = x)$.

- ▶ Construct a flow of superlevel sets of the Doob h -transform

$$S_n(t) := \{x : h_n(x, t) \geq \exp(-r(t))\}, \quad t_{n-1} < t \leq t_n,$$

generated by a map $r : [t_{n-1}, t_n] \mapsto (0, \infty)$, such that $\lim_{t \rightarrow t_n} r(t) = \infty$ to ensure $S_n(t_n) = \mathcal{C}_n$.

- ▶ Given $r(t)$, there is some $\epsilon_r \geq 0$ such that

$$\mathbb{P}(X^n(t) \in S_n(t)) \geq 1 - \epsilon_r, \quad \forall t \in (t_{n-1}, t_n].$$

If $r(t)$ is selected suitably, $S_n(t)$ captures most of the probability mass of the constrained process $X^n(t) \implies$ it is a high probability tube connecting $X(t_{n-1})$ with the terminal set \mathcal{C}_n .

How do we implement it?

- └ Modified SDEs for efficient Monte Carlo filtering
 - └ Barrier functions

Barrier functions (Erdogan *et al.*, 2025)

- ▶ Approximate high-probability tube
 - Define the constraint by way of superlevel sets of the log-likelihood

$$\mathcal{C}_n := \{x \in \mathcal{X} : \log g_n(x) > \log \zeta_n\}, \quad \text{for a threshold } \zeta_n > 0.$$

- Define the flow of compact sets

$$\hat{\mathcal{S}}_n(t) := \{x \in \mathcal{X} : \log g^n(x, t) > \log \zeta_n\}, \quad t \in [t_{n-1}, t_n],$$

with interpolated likelihoods

$$g^n(x, t_{n-1}) = g_{n-1}(x) \quad \text{and} \quad g^n(x, t_n) = g_n(x).$$

- $\hat{\mathcal{S}}_n(t)$ interpolates $\hat{\mathcal{S}}_n(t_{n-1}) = \mathcal{C}_{n-1}$ and $\hat{\mathcal{S}}_n(t_n) := \mathcal{C}_n$

- ▶ The tube $\hat{\mathcal{S}}_n(t)$ is imposed as a soft-constraint,

$$dX^n(t) = \left[a(X^n(t), t) - \mu \sigma^2(X^n(t), t) \nabla \overbrace{\log b_n(X^n(t), t)}^{\text{log-barrier}} \right] dt + \sigma(X^n(t), t) dW(t)$$

The barrier $-b_n(x, t)$ is a tractable surrogate for the Doob transform $h_n(x, t)$

Implementation of barrier functions

- ▶ We have used a shifted soft-plus function with parameters ϱ and $\kappa \geq 1$,

$$u_{\varrho, \kappa}(z) := \frac{1}{\kappa} \log(1 + \exp\{\kappa(z - \varrho)\}).$$

Parameter ϱ is a radius that determines the boundaries of the tube: for large enough κ ,

$$u_{\varrho, \kappa}(z) \approx \begin{cases} 0 & \text{when } z < \varrho \\ z - \varrho & \text{when } z > \varrho \end{cases}$$

- ▶ The barrier function is then constructed as

$$\log b_n(x, t) := u_{\varrho, \kappa}(-\log g^n(x, t)), \quad t \in [t_{n-1}, t_n].$$

Example: Gaussian observations $Y_n = H^\top X_n + U_n$, $U_n \sim N(0, \sigma_y^2 I)$

The tube is

$$\hat{S}_n(t) = \{x \in \mathcal{X} : \log g^n(x, t) \geq \log \zeta_n\} = \left\{x \in \mathcal{X} : \|y^n(t) - H^\top x\| \leq \sqrt{-\frac{\log \zeta_n}{2\sigma_y^2}}\right\} \text{ and}$$

$$\nabla \log b_n(x, t) = \frac{1}{\sigma_y^2} s \left(\kappa \left(\frac{1}{2\sigma_y^2} \|y^n(t) - H^\top x\|^2 - \varrho_n \right) \right) H(H^\top x - y^n(t)), \text{ where}$$

$\varrho_n = -\log \zeta_n$ and $s(z) = 1/(1 + e^{-z})$ is the logistic function.

- Modified SDEs for efficient Monte Carlo filtering

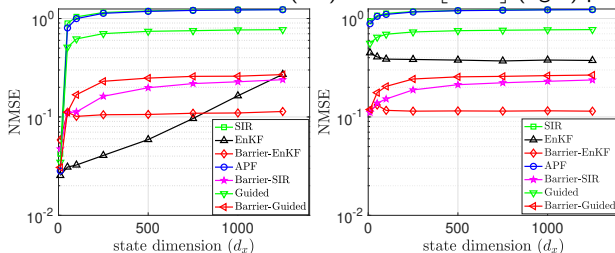
- Barrier functions

Numerics: stochastic Lorenz 96 w/ additive noise

$$dX_i(t) = [X_{i-1}(t)(X_{i+1}(t) - X_{i-2}(t)) - X_i(t) + F] dt + \sigma_x dW_i(t), \quad i = 0, \dots, d_x - 1.$$

$$Y_n = H^T X_n + U_n, \quad U_n \sim N(0, \sigma_y^2 I).$$

Monte Carlo filters with $M = 750$ (left) and $M = \lfloor 0.5d_x \rfloor$ (right) particles



state dim. d_x	1000	1250	2500	5000	7500	10,000
Barrier-ENKF, $M = 750$	0.1056	0.1129	0.1214	0.1485	0.1604	0.1801
Barrier-SIR, $M = 750$	0.2176	0.2208	0.2560	0.2705	0.2757	0.2794
Barrier-Guided, $M = 750$	0.2478	0.2562	0.2770	0.2762	0.2797	0.2910

Discussion

Discussion

- ▶ Bayesian filtering for discretely-observed diffusions (continuous-discrete state space models)
- ▶ Time-discretisation & Monte Carlo errors
- ▶ Efficient conditional sampling of SDEs
 - Feedback particle filter (Yang *et al.*, 2014)
 - Doob h -transforms (Chopin *et al.*, 2023; Pieper-Sethmacher *et al.*, 2025b,a)

are hardly practical in high-dimensional, nonlinear settings

- ▶ Constrained filters and barrier functions
 - borrowed from control theory (Margellos & Lygeros, 2011; Clark, 2021)
 - barriers are tractable surrogates of intractable Doob h -transforms
 - promising numerics for high-dimensional filtering

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